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Indeed,  $\phi_X(t) = \exp\{it^T 0 - \frac{1}{2}t^T I t\} = \exp\{-t^T t/2\} = \psi(t^T t)$ , and thus  $X$  has a spherical distribution.

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# Elliptical distributions

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The characteristic function can be written as

$$\begin{aligned}\phi_X(t) &= E(\exp\{it^T X\}) = E(\exp\{it^T(\mu + AY)\}) \\ &= \exp\{it^T \mu\} E(\exp\{i(A^T t)^T Y\}) \\ &= \exp\{it^T \mu\} \psi(t^T \Sigma t),\end{aligned}$$

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If  $A \in \mathbb{R}^{d \times d}$  is nonsingular, then we have the following relation between elliptical and spherical distributions:

$X \sim E_d(\mu, \Sigma, \psi) \Leftrightarrow A^{-1}(X - \mu) \sim S_d(\psi)$ ,  $A \in \mathbb{R}^{d \times d}$ ,  $AA^T = \Sigma$ .

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**Theorem:** (Stochastic representation of elliptical distributions)

Let  $X \in \mathbb{R}^d$  be an  $d$ -dimensional random vector.  $X \sim E_d(\mu, \Sigma, \psi)$  iff  $X \stackrel{d}{=} \mu + RAS$ , where  $S \in \mathbb{R}^k$  is a random vector uniformly distributed on the unit sphere  $\mathcal{S}^{k-1}$ ,  $R \geq 0$  is a r.v. independent of  $S$ ,  $A \in \mathbb{R}^{d \times k}$  is a constant matrix with  $\Sigma = AA^T$  and  $\mu \in \mathbb{R}^d$  is a constant vector.

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- (iii) Set  $x = \mu + rAs$ .

# Examples of elliptical distributions

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- ▶ Multivariate normal distribution

Let  $X \sim N(\mu, \Sigma)$  with  $\Sigma$  positive definite. Then for  $A \in \mathbb{R}^{d \times k}$  with  $AA^T = \Sigma$  we have  $X \stackrel{d}{=} \mu + AZ$ , where  $Z \in N_k(0, I)$ . Moreover  $Z = RS$  holds with  $S$  being uniformly distributed on the unit sphere  $\mathcal{S}^{k-1}$  and  $R^2 \sim \chi_k^2$ . Thus  $X \stackrel{d}{=} \mu + RAS$  holds and hence  $X \sim E_d(\mu, \Sigma, \psi)$  with  $\psi(x) = \exp\{-x/2\}$ .

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- ▶ Multivariate normal variance mixtures

Let  $Z \sim N_d(0, I)$ . Then  $Z$  has a spherical distribution with stochastic representation  $Z \stackrel{d}{=} VS$  with  $V^2 = \|Z\|^2 \sim \chi_d^2$ . Let  $X = \mu + WAZ$  be a variance normal mixture distribution. Then we get  $X \stackrel{d}{=} \mu + VWAS$  with  $V^2 \sim \chi_d^2$  and  $VW$  is a nonnegative r.v. independent of  $S$ . Thus  $X$  is elliptically distributed with  $R = VW$ .

# Properties of elliptical distributions

## Theorem:

Let  $X \sim E_k(\mu, \Sigma, \psi)$ .  $X$  has the following properties:

- ▶ Linear combinations

For  $B \in \mathbb{R}^{k \times d}$  and  $b \in \mathbb{R}^k$  we have:

$$BX + b \in E_k(B\mu + b, B\Sigma B^T, \psi).$$

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- ▶ Marginal distributions

Set  $X^T = \left( X^{(1)T}, X^{(2)T} \right)$  for  $X^{(1)T} = (X_1, X_2, \dots, X_n)^T$  and

$X^{(2)T} = (X_{n+1}, X_{n+2}, \dots, X_k)^T$  and analogously set

$\mu^T = \left( \mu^{(1)T}, \mu^{(2)T} \right)$  as well as  $\Sigma = \begin{pmatrix} \Sigma^{(1,1)} & \Sigma^{(1,2)} \\ \Sigma^{(2,1)} & \Sigma^{(2,2)} \end{pmatrix}$ . Then

$X_1 \sim E_n\left(\mu^{(1)}, \Sigma^{(1,1)}, \psi\right)$  and  $X_2 \sim E_{k-n}\left(\mu^{(2)}, \Sigma^{(2,2)}, \psi\right)$ .

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- ▶ Conditional distributions

Assume that  $\Sigma$  is nonsingular. Then

$$X^{(2)} \mid X^{(1)} = x^{(1)} \sim E_{d-k} \left( \mu^{(2,1)}, \Sigma^{(22,1)}, \tilde{\psi} \right) \text{ where}$$

$$\mu^{(2,1)} = \mu^{(2)} + \Sigma^{(2,1)} \left( \Sigma^{(1,1)} \right)^{-1} \left( x^{(1)} - \mu^{(1)} \right) \text{ and}$$

$$\Sigma^{(22,1)} = \Sigma^{(2,2)} - \Sigma^{(2,1)} \left( \Sigma^{(1,1)} \right)^{-1} \Sigma^{(1,2)}.$$

Typically  $\tilde{\psi}$  is a different characteristic generator than the original  $\psi$  (see Fang, Katz and Ng 1987).

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- ▶ Quadratic forms

If  $\Sigma$  is nonsingular, then  $D^2 = (X - \mu)^T \Sigma^{-1} (X - \mu) \sim R^2$ , where  $R$  is the nonnegative r.v. in the stochastic representation  $Y = RS$  of the spherical distribution  $Y$  with  $S \sim U\left(\mathcal{S}^{(d-1)}\right)$  and  $X = \mu + AY$ .  
The random variable  $D$  is called *Mahalanobis distance*.

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- ▶ Convolutions

Let  $X \sim E_k(\mu, \Sigma, \psi)$  and  $Y \sim E_k(\tilde{\mu}, \Sigma, \tilde{\psi})$  be two independent random vectors. Then  $X + Y \sim E_k(\mu + \tilde{\mu}, \Sigma, \bar{\psi})$  where  $\bar{\psi} = \psi \tilde{\psi}$ .

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**Important:**  $X \sim E_k(\mu, I_k, \psi)$  does not imply that the components of  $X$  are independent. The components of  $X$  are independent iff  $X$  is multivariate normally distributed with the unit matrix as a covariance matrix.

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(C4) Monotonicity:

$$\forall X_1, X_2 \in M \text{ the implication } X_1 \stackrel{\text{a.s.}}{\leq} X_2 \implies \rho(X_1) \leq \rho(X_2) \text{ holds.}$$

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Consider the property:

(C5) Convexity:

$$\forall X_1, X_2 \in M, \forall \lambda \in [0, 1]$$

$$\rho(\lambda X_1 + (1 - \lambda)X_2) \leq \lambda \rho(X_1) + (1 - \lambda)\rho(X_2) \text{ holds.}$$

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Let the probability measure  $P$  be defined by some continuous or discrete probability distribution  $F$ .

$VaR_\alpha(F) = F^{\leftarrow}(\alpha)$  has the properties (C1), (C3) and (C4), whereas the subadditivity (C2) is not fulfilled in general.

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(C1),(C3), (C4) follow from  $CVaR_\alpha(F) = \frac{1}{1-\alpha} \int_\alpha^1 Var_p(F) dp$ .

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To show (C2) observe that for a sequence of i.i.d. r.v.  $L_1, L_2, \dots, L_n$  with order statistics  $L_{1,n} \geq L_{2,n} \geq \dots \geq L_{n,n}$  and for any  $m \in \{1, 2, \dots, n\}$

$$\sum_{i=1}^m L_{i,n} = \sup\{L_{i_1} + L_{i_2} + \dots + L_{i_m} : 1 \leq i_1 < \dots < i_m \leq n\} \text{ holds.}$$



# The mean-risk portfolio optimization model

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Consider a portfolio of  $d$  risky assets and the random vector  $X = (X_1, X_2, \dots, X_d)^T$  of their returns. Let  $E(X) = \mu$ .

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For a risk measure  $\rho$  the *mean- $\rho$  portfolio optimization model* is:

$$\min_{w \in \mathcal{P}_m} \rho(Z(w)) \tag{1}$$

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If  $\rho = \text{VaR}_\alpha$ ,  $\alpha \in (0, 1)$  we get the *mean-VaR pf. optimization model*

$$\min_{w \in \mathcal{P}_m} \text{VaR}_\alpha(Z(w)).$$

Question: What is the relationship between these three portfolio optimization models?