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## Simulation of t-copulas

## Simulation of $t$-copulas

Algorithm: for the generation of a random vector $U=\left(U_{1}, U_{2}, \ldots, U_{d}\right)$ whose distribution function is the copula $C_{\nu, R}^{t}, R$ positive definite with all ones on the main diagonal, $\nu \in \mathbb{N}$.

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The generator $\varphi(t)=\left(t^{-\theta}-1\right) / \theta, \theta>0$ yields the Clayton copula $C_{\theta}^{C l}$.

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For $X \sim \operatorname{Gamma}(1 / \theta, 1)$ with d.f. $f_{X}(x)=\left(x^{1 / \theta-1} e^{-x}\right) / \Gamma(1 / \theta)$ we have:
$E\left(e^{-s X}\right)=\int_{0}^{\infty} e^{-s x} \frac{1}{\Gamma(1 / \theta)} x^{1 / \theta-1} e^{-x} d x=(s+1)^{-1 / \theta}=\tilde{\varphi}^{-1}(s)$.

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The simulation of $Z \sim S T(\alpha, \beta, 1,0)$ is not straightforward (see Nolan 2002).

For $\alpha \neq 1$ we get: $X=\delta+\gamma Z \sim \operatorname{St}(\alpha, \beta, \gamma, \delta)$.
The case $\alpha=1$ is more complicated.
Alternative approach:
Let $\theta \geq 1$ and $\bar{F}(x)=1-F(x)=\exp \left(-x^{1 / \theta}\right)$ for $x \geq 0$. Let $V \sim U(0,1)$ and let $S$ be a r.v. independent from $V$ with density function $h(s)=(1-1 / \theta+s / \theta) \exp (-s)$ for $s \geq 0$.
Set $\left(Z_{1}, Z_{2}\right)^{T}:=\left(V S^{\theta},(1-V) S^{\theta}\right)^{T}$.

## Simulation of the Gumbel copula ( $\theta \geq 1$ )

Let $X$ be a positive stable r.v., $X \sim \operatorname{St}(1 / \theta, 1, \gamma, 0)$ with
$\gamma=(\cos (\pi /(2 \theta)))^{\theta}>0\left(\right.$ and $\left.\alpha=\frac{1}{\theta}, \beta=1, \delta=0\right)$
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The distribution function of $\left(\bar{F}\left(Z_{1}\right), \bar{F}\left(Z_{2}\right)\right)^{T}$ is $C_{\theta}^{G u}$. Convince yourself!

## Simulation of the Gumbel copula $(\theta \geq 1)$ (contd.)

## Simulation of the Gumbel copula ( $\theta \geq 1$ ) (contd.)

Algorithm to generate a random vector $U=\left(U_{1}, U_{2}, \ldots, U_{d}\right)$ with the Gumbel copula $C_{\theta}^{G u}$ as distribution function.

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- Simulate two i.i.d. r.v. $V_{1}, V_{2} \sim U(0,1)$.


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Algorithm to generate a random vector $U=\left(U_{1}, U_{2}, \ldots, U_{d}\right)$ with the Gumbel copula $C_{\theta}^{G u}$ as distribution function. Input: The dimension $d \in \mathbb{N}$, the parameter $\theta \geq 1$.

- Simulate two i.i.d. r.v. $V_{1}, V_{2} \sim U(0,1)$.
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- Set $S:=I_{V_{2} \leq 1 / \theta} W_{1}+I_{V_{2}>1 / \theta} W_{2}$.
- Set $\left(Z_{1}, Z_{2}\right):=\left(V_{1} S^{\theta},\left(1-V_{1}\right) S^{\theta}\right)$.
- The distribution function of $U=\left(\exp \left(-Z_{1}^{1 / \theta}\right), \exp \left(-Z_{2}^{1 / \theta}\right)\right)^{T}$ is $C_{\theta}^{G u}$.


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Output: U

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Question 2: What are the parameters of the prespecified family of copulas used for the modelling?

Parameter estimation for $C_{R}^{G a}, C_{\nu, R}^{t}, C_{\theta}^{C l}$ and $C_{\theta}^{G u}$

$$
C_{R}^{G a}=\phi_{R}^{d}\left(\phi^{-1}\left(u_{1}\right), \ldots, \phi^{-1}\left(u_{d}\right)\right)
$$

$$
R_{i j}=\sin \left(\pi\left(\rho_{\tau}\right)_{i j} / 2\right)
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C_{\theta}^{C I}(u)=\left(u_{1}^{-\theta}+\ldots+u_{d}^{-\theta}-d+1\right)^{-1 / \theta} & \theta=2\left(\rho_{\tau}\right)_{i j} /\left(1-\left(\rho_{\tau}\right)_{i j}\right)
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\begin{aligned}
\left(\rho_{\tau}\right)_{i j} & =\rho_{\tau}\left(X_{k, i}, X_{k, j}\right) \\
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Standard empirical estimator of Kendalls Tau:
$\widehat{\rho}_{i j}=\binom{n}{2}^{-1} \sum_{1 \leq k<I \leq n} \operatorname{sign}\left(\left(X_{k, i}-X_{l, i}\right)\left(X_{k, j}-X_{l, j}\right)\right)$.

Calibration of the correlation matrix for Gaussian and t-copulas

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Eigenvalue approach (Rousseeuw and Molenberghs 1993)

- Compute the spectral decomposition $\hat{R}=\Gamma \Lambda \Gamma^{\top}$ of $\hat{R}$, where $\Lambda$ is a diagonal matrix, containing the eigenvalues of $\hat{R}$ on the diagonal, and $\Gamma$ is an orthogonal matrix with the eigenvectors of $\hat{R}$ in its columns.


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- Set $R^{*}:=D \tilde{R} D$ where $D$ is a diagonal matrix with

$$
D_{k, k}=1 / \sqrt{\tilde{R}_{k, k}} .
$$

Estimation of the number of the degrees of freedom $\nu$ for $t$-copulas

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2. Generate a pseudo-sample of the copula

$$
\hat{U}_{k}=\left(\hat{U}_{k, 1}, \hat{U}_{k, 2}, \ldots, \hat{U}_{k, d}\right):=\left(\hat{F}_{1}\left(X_{k, 1}\right), \ldots, \hat{F}_{d}\left(X_{k, d}\right)\right),
$$

for $k=1,2, \ldots, n$ (see Genest und Rivest 1993).

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- a non-parametric estimation method;
$\hat{F}_{i}$ is the empirical distribution function $\hat{F}_{i}(x)=\frac{1}{n+1} \sum_{t=1}^{n} I_{\left\{X_{t, i} \leq x\right\}}$, $1 \leq i \leq d$.

Estimation of the number of the degrees of freedom $\nu$ for $t$-copulas (contd.)

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$$
L\left(\xi ; \hat{U}_{1}, \hat{U}_{2}, \ldots, \hat{U}_{n}\right)=\Pi_{k=1}^{n} c_{\xi, R}^{t}\left(\hat{U}_{k}\right)
$$

and $c_{\xi, R}^{t}$ is the density of the $t$-copula $C_{\xi, R}^{t}$.
This implies

$$
\sum_{k=1}^{n} \ln g_{\xi, R}\left(t_{\xi}^{-1}\left(\hat{U}_{k, 1}\right), \ldots, t_{\xi}^{-1}\left(\hat{U}_{k, d}\right)\right)-\sum_{k=1}^{n} \sum_{j=1}^{d} \ln g_{\xi}\left(t_{\xi}^{-1}\left(\hat{U}_{k, j}\right)\right)
$$

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This implies

$$
\begin{gathered}
\ln L\left(\xi ; \hat{U}_{1}, \hat{U}_{2}, \ldots, \hat{U}_{n}\right)= \\
\sum_{k=1}^{n} \ln g_{\xi, R}\left(t_{\xi}^{-1}\left(\hat{U}_{k, 1}\right), \ldots, t_{\xi}^{-1}\left(\hat{U}_{k, d}\right)\right)-\sum_{k=1}^{n} \sum_{j=1}^{d} \ln g_{\xi}\left(t_{\xi}^{-1}\left(\hat{U}_{k, j}\right)\right),
\end{gathered}
$$

where $g_{\xi, R}$ is the cumulative density function of a $d$-dimensional standard $t$-distribution with $\xi$ degrees of freedom and correlation matrix $R$, and $g_{\xi}$ is the density function of a univariate standard $t$-distribution with $\xi$ degrees of freedom.

