

Tail dependence coefficients of elliptical copulas

Tail dependence coefficients of elliptical copulas

Theorem: Let $(X_1, X_2)^T$ be a normally distributed random vector. Then $\lambda_U(X_1, X_2) = \lambda_L(X_1, X_2) = 0$ holds.

Tail dependence coefficients of elliptical copulas

Theorem: Let $(X_1, X_2)^T$ be a normally distributed random vector. Then $\lambda_U(X_1, X_2) = \lambda_L(X_1, X_2) = 0$ holds.

Corollary: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and let C_ρ^{Ga} be a Gaussian copula, where ρ is the linear correlation coefficient of X_1 and X_2 . The $\lambda_U(X_1, X_2) = \lambda_L(X_1, X_2) = 0$ holds.

Tail dependence coefficients of elliptical copulas

Theorem: Let $(X_1, X_2)^T$ be a normally distributed random vector. Then $\lambda_U(X_1, X_2) = \lambda_L(X_1, X_2) = 0$ holds.

Corollary: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and let C_ρ^{Ga} be a Gaussian copula, where ρ is the linear correlation coefficient of X_1 and X_2 . The $\lambda_U(X_1, X_2) = \lambda_L(X_1, X_2) = 0$ holds.

Theorem: Let $(X_1, X_2)^T \sim t_2(0, \nu, R)$ be a random vector with a t -distribution and ν degrees of freedom, expectation 0 and linear correlation matrix R . For $R_{12} > -1$ we have

$$\lambda_U(X_1, X_2) = \lambda_L(X_1, X_2) = 2\bar{t}_{\nu+1} \left(\sqrt{\nu+1} \frac{\sqrt{1-R_{12}}}{\sqrt{1+R_{12}}} \right)$$

Tail dependence coefficients of elliptical copulas

Theorem: Let $(X_1, X_2)^T$ be a normally distributed random vector. Then $\lambda_U(X_1, X_2) = \lambda_L(X_1, X_2) = 0$ holds.

Corollary: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and let C_ρ^{Ga} be a Gaussian copula, where ρ is the linear correlation coefficient of X_1 and X_2 . The $\lambda_U(X_1, X_2) = \lambda_L(X_1, X_2) = 0$ holds.

Theorem: Let $(X_1, X_2)^T \sim t_2(0, \nu, R)$ be a random vector with a t -distribution and ν degrees of freedom, expectation 0 and linear correlation matrix R . For $R_{12} > -1$ we have

$$\lambda_U(X_1, X_2) = \lambda_L(X_1, X_2) = 2\bar{t}_{\nu+1} \left(\sqrt{\nu+1} \frac{\sqrt{1-R_{12}}}{\sqrt{1+R_{12}}} \right)$$

The proof is similar to the proof of the analogous theorem about the Gaussian copulas.

Hint:

$$X_2|X_1 = x \sim \left(\frac{\nu+1}{\nu+x^2} \right)^{1/2} \frac{X_2 - \rho x}{\sqrt{1-\rho^2}} \sim t_{\nu+1}$$

Tail dependence (contd.) and rank correlation of elliptical copulas

Tail dependence (contd.) and rank correlation of elliptical copulas

Corollary: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and a t -copula $C_{\nu, R}^t$ with ν degrees of freedom and correlation matrix R . Then we have

$$\lambda_U(X_1, X_2) = \lambda_L(X_1, X_2) = 2\bar{t}_{\nu+1} \left(\sqrt{\nu+1} \frac{\sqrt{1-R_{12}}}{\sqrt{1+R_{12}}} \right).$$

Tail dependence (contd.) and rank correlation of elliptical copulas

Corollary: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and a t -copula $C_{\nu, R}^t$ with ν degrees of freedom and correlation matrix R . Then we have

$$\lambda_U(X_1, X_2) = \lambda_L(X_1, X_2) = 2\bar{t}_{\nu+1} \left(\sqrt{\nu+1} \frac{\sqrt{1-R_{12}}}{\sqrt{1+R_{12}}} \right).$$

Theorem: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and a Gaussian copula C_{ρ}^{Ga} , where ρ is the linear correlation coefficient of X_1 and X_2 . Then we have $\rho_{\tau}(X_1, X_2) = \frac{2}{\pi} \arcsin \rho$ and $\rho_S(X_1, X_2) = \frac{6}{\pi} \arcsin \frac{\rho}{2}$.

Tail dependence (contd.) and rank correlation of elliptical copulas

Corollary: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and a t -copula $C_{\nu, R}^t$ with ν degrees of freedom and correlation matrix R . Then we have

$$\lambda_U(X_1, X_2) = \lambda_L(X_1, X_2) = 2\bar{t}_{\nu+1} \left(\sqrt{\nu+1} \frac{\sqrt{1-R_{12}}}{\sqrt{1+R_{12}}} \right).$$

Theorem: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and a Gaussian copula C_{ρ}^{Ga} , where ρ is the linear correlation coefficient of X_1 and X_2 . Then we have $\rho_{\tau}(X_1, X_2) = \frac{2}{\pi} \arcsin \rho$ and $\rho_S(X_1, X_2) = \frac{6}{\pi} \arcsin \frac{\rho}{2}$.

Theorem: Let $X \sim E_d(\mu, \Sigma, \psi)$ be an elliptically distributed random vector with continuous marginal distributions. Then the following holds $\rho_{\tau}(X_i, X_j) = \frac{2}{\pi} \arcsin R_{ij}$, with $R_{ij} = \frac{\Sigma_{ij}}{\sqrt{\Sigma_{ii}\Sigma_{jj}}}$ for $i, j = 1, 2, \dots, d$.

Tail dependence (contd.) and rank correlation of elliptical copulas

Corollary: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and a t -copula $C_{\nu, R}^t$ with ν degrees of freedom and correlation matrix R . Then we have

$$\lambda_U(X_1, X_2) = \lambda_L(X_1, X_2) = 2\bar{t}_{\nu+1} \left(\sqrt{\nu+1} \frac{\sqrt{1-R_{12}}}{\sqrt{1+R_{12}}} \right).$$

Theorem: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and a Gaussian copula C_ρ^{Ga} , where ρ is the linear correlation coefficient of X_1 and X_2 . Then we have $\rho_\tau(X_1, X_2) = \frac{2}{\pi} \arcsin \rho$ and $\rho_S(X_1, X_2) = \frac{6}{\pi} \arcsin \frac{\rho}{2}$.

Theorem: Let $X \sim E_d(\mu, \Sigma, \psi)$ be an elliptically distributed random vector with continuous marginal distributions. Then the following holds $\rho_\tau(X_i, X_j) = \frac{2}{\pi} \arcsin R_{ij}$, with $R_{ij} = \frac{\Sigma_{ij}}{\sqrt{\Sigma_{ii}\Sigma_{jj}}}$ for $i, j = 1, 2, \dots, d$.

Corollary: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and an elliptical copula $C_{\mu, \Sigma, \psi}^E$. Then we have $\rho_\tau(X_1, X_2) = \frac{2}{\pi} \arcsin R_{12}$, with $R_{12} = \frac{\Sigma_{12}}{\sqrt{\Sigma_{11}\Sigma_{22}}}$.

Tail dependence (contd.) and rank correlation of elliptical copulas

Corollary: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and a t -copula $C_{\nu, R}^t$ with ν degrees of freedom and correlation matrix R . Then we have

$$\lambda_U(X_1, X_2) = \lambda_L(X_1, X_2) = 2\bar{t}_{\nu+1} \left(\sqrt{\nu+1} \frac{\sqrt{1-R_{12}}}{\sqrt{1+R_{12}}} \right).$$

Theorem: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and a Gaussian copula C_{ρ}^{Ga} , where ρ is the linear correlation coefficient of X_1 and X_2 . Then we have $\rho_{\tau}(X_1, X_2) = \frac{2}{\pi} \arcsin \rho$ and $\rho_S(X_1, X_2) = \frac{6}{\pi} \arcsin \frac{\rho}{2}$.

Theorem: Let $X \sim E_d(\mu, \Sigma, \psi)$ be an elliptically distributed random vector with continuous marginal distributions. Then the following holds $\rho_{\tau}(X_i, X_j) = \frac{2}{\pi} \arcsin R_{ij}$, with $R_{ij} = \frac{\Sigma_{ij}}{\sqrt{\Sigma_{ii}\Sigma_{jj}}}$ for $i, j = 1, 2, \dots, d$.

Corollary: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and an elliptical copula $C_{\mu, \Sigma, \psi}^E$. Then we have $\rho_{\tau}(X_1, X_2) = \frac{2}{\pi} \arcsin R_{12}$, with $R_{12} = \frac{\Sigma_{12}}{\sqrt{\Sigma_{11}\Sigma_{22}}}$.

See McNeil et al. (2005) for a proof of the three last results.

Archimedean copulas

Archimedean copulas

Disadvantages of elliptical copulas:

- ▶ no closed form representation in general,
- ▶ radial symmetry

Archimedean copulas

Disadvantages of elliptical copulas:

- ▶ no closed form representation in general,
- ▶ radial symmetry

Definition: Let $\phi: [0, 1] \rightarrow [0, +\infty]$ be a continuous, strictly monotone decreasing function with $\phi(1) = 0$. The pseudo-inverse function $\phi^{[-1]}: [0, \infty] \rightarrow [0, 1]$ of ϕ is defined by

$$\phi^{[-1]}(t) = \begin{cases} \phi^{-1}(t) & 0 \leq t \leq \phi(0) \\ 0 & \phi(0) \leq t \leq \infty \end{cases}$$

Archimedean copulas

Disadvantages of elliptical copulas:

- ▶ no closed form representation in general,
- ▶ radial symmetry

Definition: Let $\phi: [0, 1] \rightarrow [0, +\infty]$ be a continuous, strictly monotone decreasing function with $\phi(1) = 0$. The pseudo-inverse function $\phi^{[-1]}: [0, \infty] \rightarrow [0, 1]$ of ϕ is defined by

$$\phi^{[-1]}(t) = \begin{cases} \phi^{-1}(t) & 0 \leq t \leq \phi(0) \\ 0 & \phi(0) \leq t \leq \infty \end{cases}$$

$\phi^{[-1]}$ is continuous and monotone decreasing on $[0, \infty]$, strictly monotone decreasing on $[0, \phi(0)]$ and $\phi^{[-1]}(\phi(u)) = u$ for $u \in [0, 1]$ holds. Moreover

$$\phi(\phi^{[-1]}(t)) = \begin{cases} t & 0 \leq t \leq \phi(0) \\ \phi(0) & \phi(0) \leq t \leq +\infty \end{cases}$$

Archimedean copulas

Disadvantages of elliptical copulas:

- ▶ no closed form representation in general,
- ▶ radial symmetry

Definition: Let $\phi: [0, 1] \rightarrow [0, +\infty]$ be a continuous, strictly monotone decreasing function with $\phi(1) = 0$. The pseudo-inverse function $\phi^{[-1]}: [0, \infty] \rightarrow [0, 1]$ of ϕ is defined by

$$\phi^{[-1]}(t) = \begin{cases} \phi^{-1}(t) & 0 \leq t \leq \phi(0) \\ 0 & \phi(0) \leq t \leq \infty \end{cases}$$

$\phi^{[-1]}$ is continuous and monotone decreasing on $[0, \infty]$, strictly monotone decreasing on $[0, \phi(0)]$ and $\phi^{[-1]}(\phi(u)) = u$ for $u \in [0, 1]$ holds. Moreover

$$\phi(\phi^{[-1]}(t)) = \begin{cases} t & 0 \leq t \leq \phi(0) \\ \phi(0) & \phi(0) \leq t \leq +\infty \end{cases}$$

If $\phi(0) = +\infty$, then $\phi^{[-1]} = \phi^{-1}$.

Archimedean copulas (contd.)

Archimedean copulas (contd.)

Theorem: Let $\phi: [0, 1] \rightarrow [0, +\infty]$ be a continuous, strictly monotone decreasing function with $\phi(1) = 0$ and let $\phi^{[-1]}$ be the pseudo-inverse function of ϕ . The function $C: [0, 1]^2 \rightarrow [0, 1]$, with $C(u_1, u_2) = \phi^{[-1]}(\phi(u_1) + \phi(u_2))$ is a copula iff ϕ is convex.

Archimedean copulas (contd.)

Theorem: Let $\phi: [0, 1] \rightarrow [0, +\infty]$ be a continuous, strictly monotone decreasing function with $\phi(1) = 0$ and let $\phi^{[-1]}$ be the pseudo-inverse function of ϕ . The function $C: [0, 1]^2 \rightarrow [0, 1]$, with $C(u_1, u_2) = \phi^{[-1]}(\phi(u_1) + \phi(u_2))$ is a copula iff ϕ is convex. A copula C generated as above is called an *Archimedean copula* with *generator* ϕ .

Archimedean copulas (contd.)

Theorem: Let $\phi: [0, 1] \rightarrow [0, +\infty]$ be a continuous, strictly monotone decreasing function with $\phi(1) = 0$ and let $\phi^{[-1]}$ be the pseudo-inverse function of ϕ . The function $C: [0, 1]^2 \rightarrow [0, 1]$, with

$C(u_1, u_2) = \phi^{[-1]}(\phi(u_1) + \phi(u_2))$ is a copula iff ϕ is convex.

A copula C generated as above is called an *Archimedean copula* with *generator* ϕ .

If $\phi(0) = +\infty$, then $\phi^{[-1]} = \phi^{-1}$ and $C(u_1, u_2) = \phi^{-1}(\phi(u_1) + \phi(u_2))$.

Archimedean copulas (contd.)

Theorem: Let $\phi: [0, 1] \rightarrow [0, +\infty]$ be a continuous, strictly monotone decreasing function with $\phi(1) = 0$ and let $\phi^{[-1]}$ be the pseudo-inverse function of ϕ . The function $C: [0, 1]^2 \rightarrow [0, 1]$, with $C(u_1, u_2) = \phi^{[-1]}(\phi(u_1) + \phi(u_2))$ is a copula iff ϕ is convex.

A copula C generated as above is called an *Archimedean copula* with *generator* ϕ .

If $\phi(0) = +\infty$, then $\phi^{[-1]} = \phi^{-1}$ and $C(u_1, u_2) = \phi^{-1}(\phi(u_1) + \phi(u_2))$.

See Nelsen 1999 for a proof

Archimedean copulas (contd.)

Theorem: Let $\phi: [0, 1] \rightarrow [0, +\infty]$ be a continuous, strictly monotone decreasing function with $\phi(1) = 0$ and let $\phi^{[-1]}$ be the pseudo-inverse function of ϕ . The function $C: [0, 1]^2 \rightarrow [0, 1]$, with $C(u_1, u_2) = \phi^{[-1]}(\phi(u_1) + \phi(u_2))$ is a copula iff ϕ is convex.

A copula C generated as above is called an *Archimedean copula* with *generator* ϕ .

If $\phi(0) = +\infty$, then $\phi^{[-1]} = \phi^{-1}$ and $C(u_1, u_2) = \phi^{-1}(\phi(u_1) + \phi(u_2))$.

See Nelsen 1999 for a proof

Examples: Gumbel Copulas: $\phi(t) = (-\ln t)^\theta$, $\theta \geq 1$, $t \in [0, 1]$. Then $C_\theta^{Gu}(u_1, u_2) = \phi^{[-1]}(\phi(u_1) + \phi(u_2)) = \exp(-[(-\ln u_1)^\theta + (-\ln u_2)^\theta]^{1/\theta})$ is the Gumbel copula with parameter θ .

Archimedean copulas (contd.)

Theorem: Let $\phi: [0, 1] \rightarrow [0, +\infty]$ be a continuous, strictly monotone decreasing function with $\phi(1) = 0$ and let $\phi^{[-1]}$ be the pseudo-inverse function of ϕ . The function $C: [0, 1]^2 \rightarrow [0, 1]$, with $C(u_1, u_2) = \phi^{[-1]}(\phi(u_1) + \phi(u_2))$ is a copula iff ϕ is convex.

A copula C generated as above is called an *Archimedean copula* with *generator* ϕ .

If $\phi(0) = +\infty$, then $\phi^{[-1]} = \phi^{-1}$ and $C(u_1, u_2) = \phi^{-1}(\phi(u_1) + \phi(u_2))$.

See Nelsen 1999 for a proof

Examples: Gumbel Copulas: $\phi(t) = (-\ln t)^\theta$, $\theta \geq 1$, $t \in [0, 1]$. Then $C_\theta^{Gu}(u_1, u_2) = \phi^{[-1]}(\phi(u_1) + \phi(u_2)) = \exp(-[(-\ln u_1)^\theta + (-\ln u_2)^\theta]^{1/\theta})$ is the Gumbel copula with parameter θ .

For $\theta = 1$: $C_1^{Gu} = u_1 u_2$.

Archimedean copulas (contd.)

Theorem: Let $\phi: [0, 1] \rightarrow [0, +\infty]$ be a continuous, strictly monotone decreasing function with $\phi(1) = 0$ and let $\phi^{[-1]}$ be the pseudo-inverse function of ϕ . The function $C: [0, 1]^2 \rightarrow [0, 1]$, with

$C(u_1, u_2) = \phi^{[-1]}(\phi(u_1) + \phi(u_2))$ is a copula iff ϕ is convex.

A copula C generated as above is called an *Archimedean copula* with *generator* ϕ .

If $\phi(0) = +\infty$, then $\phi^{[-1]} = \phi^{-1}$ and $C(u_1, u_2) = \phi^{-1}(\phi(u_1) + \phi(u_2))$.

See Nelsen 1999 for a proof

Examples: Gumbel Copulas: $\phi(t) = (-\ln t)^\theta$, $\theta \geq 1$, $t \in [0, 1]$. Then $C_\theta^{Gu}(u_1, u_2) = \phi^{[-1]}(\phi(u_1) + \phi(u_2)) = \exp(-[(-\ln u_1)^\theta + (-\ln u_2)^\theta]^{1/\theta})$ is the Gumbel copula with parameter θ .

For $\theta = 1$: $C_1^{Gu} = u_1 u_2$. $\lim_{\theta \rightarrow \infty} C_\theta^{Gu} = M(u_1, u_2) = \min\{u_1, u_2\}$.

Archimedean copulas (contd.)

Theorem: Let $\phi: [0, 1] \rightarrow [0, +\infty]$ be a continuous, strictly monotone decreasing function with $\phi(1) = 0$ and let $\phi^{[-1]}$ be the pseudo-inverse function of ϕ . The function $C: [0, 1]^2 \rightarrow [0, 1]$, with $C(u_1, u_2) = \phi^{[-1]}(\phi(u_1) + \phi(u_2))$ is a copula iff ϕ is convex.

A copula C generated as above is called an *Archimedean copula* with *generator* ϕ .

If $\phi(0) = +\infty$, then $\phi^{[-1]} = \phi^{-1}$ and $C(u_1, u_2) = \phi^{-1}(\phi(u_1) + \phi(u_2))$.

See Nelsen 1999 for a proof

Examples: Gumbel Copulas: $\phi(t) = (-\ln t)^\theta$, $\theta \geq 1$, $t \in [0, 1]$. Then $C_\theta^{Gu}(u_1, u_2) = \phi^{[-1]}(\phi(u_1) + \phi(u_2)) = \exp(-[(-\ln u_1)^\theta + (-\ln u_2)^\theta]^{1/\theta})$ is the Gumbel copula with parameter θ .

For $\theta = 1$: $C_1^{Gu} = u_1 u_2$. $\lim_{\theta \rightarrow \infty} C_\theta^{Gu} = M(u_1, u_2) = \min\{u_1, u_2\}$.

The Gumbel Copulas have an upper tail dependence.

Archimedean copulas (contd.)

Theorem: Let $\phi: [0, 1] \rightarrow [0, +\infty]$ be a continuous, strictly monotone decreasing function with $\phi(1) = 0$ and let $\phi^{[-1]}$ be the pseudo-inverse function of ϕ . The function $C: [0, 1]^2 \rightarrow [0, 1]$, with $C(u_1, u_2) = \phi^{[-1]}(\phi(u_1) + \phi(u_2))$ is a copula iff ϕ is convex.

A copula C generated as above is called an *Archimedean copula* with *generator* ϕ .

If $\phi(0) = +\infty$, then $\phi^{[-1]} = \phi^{-1}$ and $C(u_1, u_2) = \phi^{-1}(\phi(u_1) + \phi(u_2))$.

See Nelsen 1999 for a proof

Examples: Gumbel Copulas: $\phi(t) = (-\ln t)^\theta$, $\theta \geq 1$, $t \in [0, 1]$. Then $C_\theta^{Gu}(u_1, u_2) = \phi^{[-1]}(\phi(u_1) + \phi(u_2)) = \exp(-[(-\ln u_1)^\theta + (-\ln u_2)^\theta]^{1/\theta})$ is the Gumbel copula with parameter θ .

For $\theta = 1$: $C_1^{Gu} = u_1 u_2$. $\lim_{\theta \rightarrow \infty} C_\theta^{Gu} = M(u_1, u_2) = \min\{u_1, u_2\}$.

The Gumbel Copulas have an upper tail dependence.

Clayton Copulas: $\phi(t) = (t^{-\theta} - 1)/\theta$, $\theta > 0$. Then

$C_\theta^{Cl}(u_1, u_2) = \phi^{[-1]}(\phi(u_1) + \phi(u_2)) = (u_1^{-\theta} + u_2^{-\theta} - 1)^{-1/\theta}$ is the Clayton copula with parameter θ .

Archimedean copulas (contd.)

Theorem: Let $\phi: [0, 1] \rightarrow [0, +\infty]$ be a continuous, strictly monotone decreasing function with $\phi(1) = 0$ and let $\phi^{[-1]}$ be the pseudo-inverse function of ϕ . The function $C: [0, 1]^2 \rightarrow [0, 1]$, with $C(u_1, u_2) = \phi^{[-1]}(\phi(u_1) + \phi(u_2))$ is a copula iff ϕ is convex.

A copula C generated as above is called an *Archimedean copula* with *generator* ϕ .

If $\phi(0) = +\infty$, then $\phi^{[-1]} = \phi^{-1}$ and $C(u_1, u_2) = \phi^{-1}(\phi(u_1) + \phi(u_2))$.

See Nelsen 1999 for a proof

Examples: Gumbel Copulas: $\phi(t) = (-\ln t)^\theta$, $\theta \geq 1$, $t \in [0, 1]$. Then $C_\theta^{Gu}(u_1, u_2) = \phi^{[-1]}(\phi(u_1) + \phi(u_2)) = \exp(-[(-\ln u_1)^\theta + (-\ln u_2)^\theta]^{1/\theta})$ is the Gumbel copula with parameter θ .

For $\theta = 1$: $C_1^{Gu} = u_1 u_2$. $\lim_{\theta \rightarrow \infty} C_\theta^{Gu} = M(u_1, u_2) = \min\{u_1, u_2\}$.

The Gumbel Copulas have an upper tail dependence.

Clayton Copulas: $\phi(t) = (t^{-\theta} - 1)/\theta$, $\theta > 0$. Then

$C_\theta^{Cl}(u_1, u_2) = \phi^{[-1]}(\phi(u_1) + \phi(u_2)) = (u_1^{-\theta} + u_2^{-\theta} - 1)^{-1/\theta}$ is the Clayton copula with parameter θ .

$\lim_{\theta \rightarrow 0} C_\theta^{Cl} = u_1 u_2$ and $\lim_{\theta \rightarrow \infty} C_\theta^{Cl} = M = \min\{u_1, u_2\}$.

Archimedean copulas (contd.)

Theorem: Let $\phi: [0, 1] \rightarrow [0, +\infty]$ be a continuous, strictly monotone decreasing function with $\phi(1) = 0$ and let $\phi^{[-1]}$ be the pseudo-inverse function of ϕ . The function $C: [0, 1]^2 \rightarrow [0, 1]$, with $C(u_1, u_2) = \phi^{[-1]}(\phi(u_1) + \phi(u_2))$ is a copula iff ϕ is convex.

A copula C generated as above is called an *Archimedean copula* with *generator* ϕ .

If $\phi(0) = +\infty$, then $\phi^{[-1]} = \phi^{-1}$ and $C(u_1, u_2) = \phi^{-1}(\phi(u_1) + \phi(u_2))$.

See Nelsen 1999 for a proof

Examples: Gumbel Copulas: $\phi(t) = (-\ln t)^\theta$, $\theta \geq 1$, $t \in [0, 1]$. Then $C_\theta^{Gu}(u_1, u_2) = \phi^{[-1]}(\phi(u_1) + \phi(u_2)) = \exp(-[(-\ln u_1)^\theta + (-\ln u_2)^\theta]^{1/\theta})$ is the Gumbel copula with parameter θ .

For $\theta = 1$: $C_1^{Gu} = u_1 u_2$. $\lim_{\theta \rightarrow \infty} C_\theta^{Gu} = M(u_1, u_2) = \min\{u_1, u_2\}$.

The Gumbel Copulas have an upper tail dependence.

Clayton Copulas: $\phi(t) = (t^{-\theta} - 1)/\theta$, $\theta > 0$. Then

$C_\theta^{Cl}(u_1, u_2) = \phi^{[-1]}(\phi(u_1) + \phi(u_2)) = (u_1^{-\theta} + u_2^{-\theta} - 1)^{-1/\theta}$ is the Clayton copula with parameter θ .

$\lim_{\theta \rightarrow 0} C_\theta^{Cl} = u_1 u_2$ and $\lim_{\theta \rightarrow \infty} C_\theta^{Cl} = M = \min\{u_1, u_2\}$.

The Clayton copulas have a lower tail dependence.

Archimedean copulas (contd.)

Archimedean copulas (contd.)

Example:

Let $\phi(t) = 1 - t$, $t \in [0, 1]$. Then $\phi^{[-1]}(t) = \max\{1 - t, 0\}$ and
 $C_\phi(u_1, u_2) := \phi^{[-1]}(\phi(u_1) + \phi(u_2)) = \max\{u_1 + u_2 - 1, 0\} = W(u_1, u_2)$.
Thus the Fréchet lower bound is an Archimedean copula.

Archimedean copulas (contd.)

Example:

Let $\phi(t) = 1 - t$, $t \in [0, 1]$. Then $\phi^{[-1]}(t) = \max\{1 - t, 0\}$ and $C_\phi(u_1, u_2) := \phi^{[-1]}(\phi(u_1) + \phi(u_2)) = \max\{u_1 + u_2 - 1, 0\} = W(u_1, u_2)$. Thus the Fréchet lower bound is an Archimedean copula.

Theorem: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and an Archimedean copula C generated by ϕ . Then $\rho_\tau(X_1, X_2) = 1 + 4 \int_0^1 \frac{\phi(t)}{\phi'(t)} dt$ holds.

Archimedean copulas (contd.)

Example:

Let $\phi(t) = 1 - t$, $t \in [0, 1]$. Then $\phi^{[-1]}(t) = \max\{1 - t, 0\}$ and $C_\phi(u_1, u_2) := \phi^{[-1]}(\phi(u_1) + \phi(u_2)) = \max\{u_1 + u_2 - 1, 0\} = W(u_1, u_2)$. Thus the Fréchet lower bound is an Archimedean copula.

Theorem: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and an Archimedean copula C generated by ϕ . Then

$$\rho_\tau(X_1, X_2) = 1 + 4 \int_0^1 \frac{\phi(t)}{\phi'(t)} dt \text{ holds.}$$

See Nelsen 1999 for a proof.

Archimedean copulas (contd.)

Example:

Let $\phi(t) = 1 - t$, $t \in [0, 1]$. Then $\phi^{[-1]}(t) = \max\{1 - t, 0\}$ and $C_\phi(u_1, u_2) := \phi^{[-1]}(\phi(u_1) + \phi(u_2)) = \max\{u_1 + u_2 - 1, 0\} = W(u_1, u_2)$. Thus the Fréchet lower bound is an Archimedean copula.

Theorem: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and an Archimedean copula C generated by ϕ . Then $\rho_\tau(X_1, X_2) = 1 + 4 \int_0^1 \frac{\phi(t)}{\phi'(t)} dt$ holds.

See Nelsen 1999 for a proof.

Example Kendalls Tau for the Gumbel copula and the Clayton copula

Archimedean copulas (contd.)

Example:

Let $\phi(t) = 1 - t$, $t \in [0, 1]$. Then $\phi^{[-1]}(t) = \max\{1 - t, 0\}$ and $C_\phi(u_1, u_2) := \phi^{[-1]}(\phi(u_1) + \phi(u_2)) = \max\{u_1 + u_2 - 1, 0\} = W(u_1, u_2)$. Thus the Fréchet lower bound is an Archimedean copula.

Theorem: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and an Archimedean copula C generated by ϕ . Then $\rho_\tau(X_1, X_2) = 1 + 4 \int_0^1 \frac{\phi(t)}{\phi'(t)} dt$ holds.

See Nelsen 1999 for a proof.

Example Kendalls Tau for the Gumbel copula and the Clayton copula

Gumbel: $\phi(t) = (\ln t)^\theta$, $\theta \geq 1$.

Archimedean copulas (contd.)

Example:

Let $\phi(t) = 1 - t$, $t \in [0, 1]$. Then $\phi^{[-1]}(t) = \max\{1 - t, 0\}$ and $C_\phi(u_1, u_2) := \phi^{[-1]}(\phi(u_1) + \phi(u_2)) = \max\{u_1 + u_2 - 1, 0\} = W(u_1, u_2)$. Thus the Fréchet lower bound is an Archimedean copula.

Theorem: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and an Archimedean copula C generated by ϕ . Then

$$\rho_\tau(X_1, X_2) = 1 + 4 \int_0^1 \frac{\phi(t)}{\phi'(t)} dt \text{ holds.}$$

See Nelsen 1999 for a proof.

Example Kendalls Tau for the Gumbel copula and the Clayton copula

Gumbel: $\phi(t) = (\ln t)^\theta$, $\theta \geq 1$.

$$\rho_\tau(\theta) = 1 + 4 \int_0^1 \frac{\phi(t)}{\phi'(t)} dt = 1 - \frac{1}{\theta}.$$

Archimedean copulas (contd.)

Example:

Let $\phi(t) = 1 - t$, $t \in [0, 1]$. Then $\phi^{[-1]}(t) = \max\{1 - t, 0\}$ and $C_\phi(u_1, u_2) := \phi^{[-1]}(\phi(u_1) + \phi(u_2)) = \max\{u_1 + u_2 - 1, 0\} = W(u_1, u_2)$. Thus the Fréchet lower bound is an Archimedean copula.

Theorem: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and an Archimedean copula C generated by ϕ . Then $\rho_\tau(X_1, X_2) = 1 + 4 \int_0^1 \frac{\phi(t)}{\phi'(t)} dt$ holds.

See Nelsen 1999 for a proof.

Example Kendalls Tau for the Gumbel copula and the Clayton copula

Gumbel: $\phi(t) = (\ln t)^\theta$, $\theta \geq 1$.

$$\rho_\tau(\theta) = 1 + 4 \int_0^1 \frac{\phi(t)}{\phi'(t)} dt = 1 - \frac{1}{\theta}.$$

Clayton: $\phi(t) = (t^{-\theta} - 1)/\theta$, $\theta > 0$.

Archimedean copulas (contd.)

Example:

Let $\phi(t) = 1 - t$, $t \in [0, 1]$. Then $\phi^{[-1]}(t) = \max\{1 - t, 0\}$ and $C_\phi(u_1, u_2) := \phi^{[-1]}(\phi(u_1) + \phi(u_2)) = \max\{u_1 + u_2 - 1, 0\} = W(u_1, u_2)$. Thus the Fréchet lower bound is an Archimedean copula.

Theorem: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and an Archimedean copula C generated by ϕ . Then $\rho_\tau(X_1, X_2) = 1 + 4 \int_0^1 \frac{\phi(t)}{\phi'(t)} dt$ holds.

See Nelsen 1999 for a proof.

Example Kendalls Tau for the Gumbel copula and the Clayton copula

Gumbel: $\phi(t) = (\ln t)^\theta$, $\theta \geq 1$.

$$\rho_\tau(\theta) = 1 + 4 \int_0^1 \frac{\phi(t)}{\phi'(t)} dt = 1 - \frac{1}{\theta}.$$

Clayton: $\phi(t) = (t^{-\theta} - 1)/\theta$, $\theta > 0$.

$$\rho_\tau(\theta) = 1 + 4 \int_0^1 \frac{\phi(t)}{\phi'(t)} dt = \frac{\theta}{\theta+2}.$$

Multivariate Archimedian copulas

Multivariate Archimedian copulas

Definition: A function $g: [0, \infty) \rightarrow [0, \infty)$ is called completely monotone iff all higher order derivatives of g exist and the following inequalities hold for $k \in \mathbb{N}_*$: $(-1)^k \left(\frac{d^k}{ds^k} g(s) \right) \Big|_{s=t} \geq 0, \forall t \in (0, \infty)$.

Multivariate Archimedean copulas

Definition: A function $g: [0, \infty) \rightarrow [0, \infty)$ is called completely monotone iff all higher order derivatives of g exist and the following inequalities hold for $k \in \mathbb{N}_*$: $(-1)^k \left(\frac{d^k}{ds^k} g(s) \right) \Big|_{s=t} \geq 0, \forall t \in (0, \infty)$.

Theorem: (Kimberling 1974)

Let $\phi: [0, 1] \rightarrow [0, \infty]$ be a continuous, strictly monotone decreasing function with $\phi(0) = \infty$ and $\phi(1) = 0$. The function $C: [0, 1]^d \rightarrow [0, 1]$, $C(u) := \phi^{-1}(\phi(u_1) + \phi(u_2) + \dots + \phi(u_d))$ is a copula for $d \geq 2$ iff ϕ^{-1} is completely monotone on $[0, \infty)$.

Multivariate Archimedean copulas

Definition: A function $g: [0, \infty) \rightarrow [0, \infty)$ is called completely monotone iff all higher order derivatives of g exist and the following inequalities hold for $k \in \mathbb{N}_*$: $(-1)^k \left(\frac{d^k}{ds^k} g(s) \right) \Big|_{s=t} \geq 0, \forall t \in (0, \infty)$.

Theorem: (Kimberling 1974)

Let $\phi: [0, 1] \rightarrow [0, \infty]$ be a continuous, strictly monotone decreasing function with $\phi(0) = \infty$ and $\phi(1) = 0$. The function $C: [0, 1]^d \rightarrow [0, 1]$, $C(u) := \phi^{-1}(\phi(u_1) + \phi(u_2) + \dots + \phi(u_d))$ is a copula for $d \geq 2$ iff ϕ^{-1} is completely monotone on $[0, \infty)$.

Lemma: A function $\psi: [0, \infty) \rightarrow [0, \infty)$ is completely monotone with $\psi(0) = 1$ iff ψ is the Laplace-Stieltjes transform of some distribution function G on $[0, \infty)$, i.e. $\psi(s) = \int_0^\infty e^{-sx} dG(x), s \geq 0$.

Multivariate Archimedean copulas (contd.)

Multivariate Archimedean copulas (contd.)

Theorem: Let G be a distribution function on $[0, \infty)$ such that $G(0) = 0$. Let ψ be the Laplace-Stieltjes transform of G , i.e. $\psi(s) = \int_0^\infty e^{-sx} dG(x)$ for $s \geq 0$. Let X be a r.v. with distribution function G and let U_1, U_2, \dots, U_d be conditionally independent r.v. for $X = x$, $x \in \mathbb{R}^+$, with conditional distribution function $F_{U_k|X=x}(u) = \exp(-x\psi^{-1}(u))$ for $u \in [0, 1]$.

Multivariate Archimedean copulas (contd.)

Theorem: Let G be a distribution function on $[0, \infty)$ such that $G(0) = 0$. Let ψ be the Laplace-Stieltjes transform of G , i.e. $\psi(s) = \int_0^\infty e^{-sx} dG(x)$ for $s \geq 0$. Let X be a r.v. with distribution function G and let U_1, U_2, \dots, U_d be conditionally independent r.v. for $X = x$, $x \in \mathbb{R}^+$, with conditional distribution function $F_{U_k|X=x}(u) = \exp(-x\psi^{-1}(u))$ for $u \in [0, 1]$.

Then

$$\text{Prob}(U_1 \leq u_1, U_2 \leq u_2, \dots, U_d \leq u_d) = \psi(\psi^{-1}(u_1) + \psi^{-1}(u_2) + \dots + \psi^{-1}(u_d))$$

and the distribution function of $U = (U_1, U_2, \dots, U_d)$ is an Archimedean copula with generator ψ^{-1} .

Multivariate Archimedean copulas (contd.)

Theorem: Let G be a distribution function on $[0, \infty)$ such that $G(0) = 0$. Let ψ be the Laplace-Stieltjes transform of G , i.e. $\psi(s) = \int_0^\infty e^{-sx} dG(x)$ for $s \geq 0$. Let X be a r.v. with distribution function G and let U_1, U_2, \dots, U_d be conditionally independent r.v. for $X = x$, $x \in \mathbb{R}^+$, with conditional distribution function $F_{U_k|X=x}(u) = \exp(-x\psi^{-1}(u))$ for $u \in [0, 1]$.

Then

$$\text{Prob}(U_1 \leq u_1, U_2 \leq u_2, \dots, U_d \leq u_d) = \psi(\psi^{-1}(u_1) + \psi^{-1}(u_2) + \dots + \psi^{-1}(u_d))$$

and the distribution function of $U = (U_1, U_2, \dots, U_d)$ is an Archimedean copula with generator ψ^{-1} .

Advantages and disadvantages of Archimedean copulas:

- ▶ can model a broader class of dependencies

Multivariate Archimedean copulas (contd.)

Theorem: Let G be a distribution function on $[0, \infty)$ such that $G(0) = 0$. Let ψ be the Laplace-Stieltjes transform of G , i.e. $\psi(s) = \int_0^\infty e^{-sx} dG(x)$ for $s \geq 0$. Let X be a r.v. with distribution function G and let U_1, U_2, \dots, U_d be conditionally independent r.v. for $X = x$, $x \in \mathbb{R}^+$, with conditional distribution function $F_{U_k|X=x}(u) = \exp(-x\psi^{-1}(u))$ for $u \in [0, 1]$.

Then

$$\text{Prob}(U_1 \leq u_1, U_2 \leq u_2, \dots, U_d \leq u_d) = \psi(\psi^{-1}(u_1) + \psi^{-1}(u_2) + \dots + \psi^{-1}(u_d))$$

and the distribution function of $U = (U_1, U_2, \dots, U_d)$ is an Archimedean copula with generator ψ^{-1} .

Advantages and disadvantages of Archimedean copulas:

- ▶ can model a broader class of dependencies
- ▶ have a closed form representation

Multivariate Archimedean copulas (contd.)

Theorem: Let G be a distribution function on $[0, \infty)$ such that $G(0) = 0$. Let ψ be the Laplace-Stieltjes transform of G , i.e. $\psi(s) = \int_0^\infty e^{-sx} dG(x)$ for $s \geq 0$. Let X be a r.v. with distribution function G and let U_1, U_2, \dots, U_d be conditionally independent r.v. for $X = x$, $x \in \mathbb{R}^+$, with conditional distribution function $F_{U_k|X=x}(u) = \exp(-x\psi^{-1}(u))$ for $u \in [0, 1]$.

Then

$$\text{Prob}(U_1 \leq u_1, U_2 \leq u_2, \dots, U_d \leq u_d) = \psi(\psi^{-1}(u_1) + \psi^{-1}(u_2) + \dots + \psi^{-1}(u_d))$$

and the distribution function of $U = (U_1, U_2, \dots, U_d)$ is an Archimedean copula with generator ψ^{-1} .

Advantages and disadvantages of Archimedean copulas:

- ▶ can model a broader class of dependencies
- ▶ have a closed form representation
- ▶ depend on a small number of parameters in general

Multivariate Archimedean copulas (contd.)

Theorem: Let G be a distribution function on $[0, \infty)$ such that $G(0) = 0$. Let ψ be the Laplace-Stieltjes transform of G , i.e. $\psi(s) = \int_0^\infty e^{-sx} dG(x)$ for $s \geq 0$. Let X be a r.v. with distribution function G and let U_1, U_2, \dots, U_d be conditionally independent r.v. for $X = x$, $x \in \mathbb{R}^+$, with conditional distribution function $F_{U_k|X=x}(u) = \exp(-x\psi^{-1}(u))$ for $u \in [0, 1]$.

Then

$$\text{Prob}(U_1 \leq u_1, U_2 \leq u_2, \dots, U_d \leq u_d) = \psi(\psi^{-1}(u_1) + \psi^{-1}(u_2) + \dots + \psi^{-1}(u_d))$$

and the distribution function of $U = (U_1, U_2, \dots, U_d)$ is an Archimedean copula with generator ψ^{-1} .

Advantages and disadvantages of Archimedean copulas:

- ▶ can model a broader class of dependencies
- ▶ have a closed form representation
- ▶ depend on a small number of parameters in general
- ▶ the generator function needs to fulfill quite restrictive technical assumptions