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Answer: Selection of a suitable family of copulas based on (a) the visual comparison of the graphical representations of the data set on one side and of known copulas on the other, and (b) the empirical tail dependence coefficients.

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Question 2: What are the parameters of the prespecified family of copulas used for the modelling?

Parameter estimation for $C_{R}^{G a}, C_{\nu, R}^{t}, C_{\theta}^{C l}$ and $C_{\theta}^{G u}$

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C_{R}^{G a}=\phi_{R}^{d}\left(\phi^{-1}\left(u_{1}\right), \ldots, \phi^{-1}\left(u_{d}\right)\right)
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C_{\theta}^{G u}(u)=\exp \left(-\left[\left(-\ln u_{1}\right)^{\theta}+\ldots+\left(-\ln u_{d}^{\theta}\right]^{1 / \theta}\right)\right. & \theta=1 /\left(1-\left(\rho_{\tau}\right)_{i j}\right)
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where

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\begin{aligned}
\left(\rho_{\tau}\right)_{i j} & =\rho_{\tau}\left(X_{k, i}, X_{k, j}\right) \\
& =P\left(\left(X_{k, i}-X_{l, i}\right)\left(X_{k, j}-X_{l, j}\right)>0\right)-P\left(\left(X_{k, i}-X_{l, i}\right)\left(X_{k, j}-X_{l, j}\right)<0\right) \\
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Standard empirical estimator of Kendalls Tau:
$\widehat{\rho}_{\tau i j}=\binom{n}{2}^{-1} \sum_{1 \leq k<I \leq n} \operatorname{sign}\left(\left(X_{k, i}-X_{l, i}\right)\left(X_{k, j}-X_{l, j}\right)\right)$.

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Eigenvalue approach (Rousseeuw and Molenberghs 1993)

- Compute the spectral decomposition $\hat{R}=\Gamma \Lambda \Gamma^{\top}$ of $\hat{R}$, where $\Lambda$ is a diagonal matrix, containing the eigenvalues of $\hat{R}$ on the diagonal, and $\Gamma$ is an orthogonal matrix with the eigenvectors of $\hat{R}$ in its columns.


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- Compute $\tilde{R}=\Gamma \tilde{\Lambda} \Gamma^{T}$. $\tilde{R}$ is symmetric and positive definite but not necessarily a correlation matrix; the diagonal elements $\hat{R}_{i i}$ might be unequal 1 .


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- Set $R^{*}:=D \tilde{R} D$ where $D$ is a diagonal matrix with

$$
D_{k, k}=1 / \sqrt{\tilde{R}_{k, k}} .
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Estimation of the number of the degrees of freedom $\nu$ for $t$-copulas

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\hat{U}_{k}=\left(\hat{U}_{k, 1}, \hat{U}_{k, 2}, \ldots, \hat{U}_{k, d}\right):=\left(\hat{F}_{1}\left(X_{k, 1}\right), \ldots, \hat{F}_{d}\left(X_{k, d}\right)\right),
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for $k=1,2, \ldots, n$ (see Genest und Rivest 1993).

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- a non-parametric estimation method;
$\hat{F}_{i}$ is the empirical distribution function $\hat{F}_{i}(x)=\frac{1}{n+1} \sum_{t=1}^{n} I_{\left\{X_{t, i} \leq x\right\}}$, $1 \leq i \leq d$.

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Maximum likelihood estimator of $\nu: \nu=\arg \max _{\xi} \ln L\left(\xi ; \hat{U}_{1}, \hat{U}_{2}, \ldots, \hat{U}_{n}\right)$

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L\left(\xi ; \hat{U}_{1}, \hat{U}_{2}, \ldots, \hat{U}_{n}\right)=\Pi_{k=1}^{n} c_{\xi, R}^{t}\left(\hat{U}_{k}\right)
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and $c_{\xi, R}^{t}$ is the density of the $t$-copula $C_{\xi, R}^{t}$.
This implies

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\begin{gathered}
\ln L\left(\xi ; \hat{U}_{1}, \hat{U}_{2}, \ldots, \hat{U}_{n}\right)= \\
\sum_{k=1}^{n} \ln g_{\xi, R}\left(t_{\xi}^{-1}\left(\hat{U}_{k, 1}\right), \ldots, t_{\xi}^{-1}\left(\hat{U}_{k, d}\right)\right)-\sum_{k=1}^{n} \sum_{j=1}^{d} \ln g_{\xi}\left(t_{\xi}^{-1}\left(\hat{U}_{k, j}\right)\right),
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\end{gathered}
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where $g_{\xi, R}$ is the density of a $d$-dimensional standard $t$-distribution with distribution function $t_{\xi, R}^{d}$ and $g_{\xi}$ is the density of a univariate standard $t$-distribution with $\xi$ degrees of freedom.

