

## Characterization of the domain of attraction

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Remark: Let  $F \in DA(G_\alpha)$  for  $\alpha \in (0, 2)$ . Then  $E(|X|^\delta) < \infty$  for  $\delta < \alpha$  and  $E(|X|^\delta) = \infty$  for  $\delta > \alpha$  hold.

## Limit distributions of normalized and centered maxima

Let  $(X_k)$ ,  $k \in \mathbb{N}$ , be non-degenerate i.i.d. r.v. with distribution function  $F$ .

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Consider  $\lim_{n \rightarrow \infty} P(a_n^{-1}(M_n - b_n) \leq x) = \lim_{n \rightarrow \infty} P(M_n \leq u_n)$ , where  $u_n = a_n x + b_n$ ,  $\forall n \in \mathbb{N}$ .



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**Theorem:** (Poisson Approximation)

Let  $\tau \in [0, \infty]$  and a sequence of reals  $u_n \in \mathbb{R}$ . Then the following holds

$$\lim_{n \rightarrow \infty} n\bar{F}(u_n) = \tau \iff \lim_{n \rightarrow \infty} P(M_n \leq u_n) = \exp\{-\tau\}.$$

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**Remark:** The convergence to types theorem implies that  $H$  and  $\tilde{H}$  are of the same type, if

$$\lim_{n \rightarrow \infty} a_n^{-1}(M_n - b_n) = H \text{ and } \lim_{n \rightarrow \infty} \tilde{a}_n^{-1}(M_n - \tilde{b}_n) = \tilde{H}.$$

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**Definition:** A non-degenerate r.v.  $X$  is called *max-stable* iff for any  $n \geq 2$   $\max\{X_1, X_2, \dots, X_n\} \stackrel{d}{=} a_n X + b_n$  for independent copies  $X_1, X_2, \dots, X_n$  of  $X$  and appropriate constants  $a_n > 0$  and  $b_n \in \mathbb{R}$ .

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**Theorem:** (Proof in McNeil, Frey und Embrechts, 2005.)

The class of max-stable distributions coincides with the class of non-degenerate limit distributions of normalized and centered maxima of i.i.d. r.v.

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**Theorem:** (Fischer-Tippet Theorem, Proof in Resnick 1987, page 9-11)  
Let  $(X_k)$  be a sequence of i.i.d. r.v.. If the constants  $a_n, b_n \in \mathbb{R}$ ,  $a_n > 0$ , and a non-degenerate distribution  $H$  exist, such that  $\lim_{n \rightarrow \infty} a_n^{-1}(M_n - b_n) = H$ , then  $H$  is of the same type as one of the following three distributions:

$$\begin{array}{ll} \text{Fréchet} & \Phi_\alpha(x) = \begin{cases} 0 & x \leq 0 \\ \exp\{-x^{-\alpha}\} & x > 0 \end{cases} & \alpha > 0 \\ \text{Weibull} & \Psi_\alpha(x) = \begin{cases} \exp\{-(-x)^\alpha\} & x \leq 0 \\ 1 & x > 0 \end{cases} & \alpha > 0 \\ \text{Gumbel} & \Lambda(x) = \exp\{-e^{-x}\} & x \in \mathbb{R} \end{array}$$

## Extreme value distributions

The distributions  $\Phi_\alpha$ ,  $\Psi_\alpha$  and  $\Lambda$  are called *standard extreme value distributions (standard evd)*. The distributions which are of the same type as the standard evd are called *extreme value distributions (evd)*.



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**Definition:** We say that the r.v.  $X$  (or the corresponding distribution) belongs to the *maximum domain of attraction* of the evd  $H$  iff there exist constants  $a_n > 0$  and  $b_n \in \mathbb{R}$  such that  $\lim_{n \rightarrow \infty} a_n^{-1}(M_n - b_n) = H$  holds. Notation:  $X \in MDA(H)$  ( $F \in MDA(H)$ ).

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**Theorem:** (Characterisation of MDA, proof is left as an exercise)  
 $F \in MDA(H)$  with normalizing and centering constants  $a_n > 0$  and  $b_n \in \mathbb{R}$  holds, iff

$$\lim_{n \rightarrow \infty} n\bar{F}(a_n x + b_n) = -\ln H(x), \forall x \in \mathbb{R},$$

where  $-\ln H(x)$  is replaced by  $\infty$  if  $H(x) = 0$ .

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There exist distributions which do not belong to the MDA of any evd!

**Example:** (The Poisson distribution)

Let  $X \sim P(\lambda)$ , i.e.  $P(X = k) = e^{-\lambda} \lambda^k / k!$ ,  $k \in \mathbb{N}_0$ ,  $\lambda > 0$ . Show that there exist no evd  $Z$  such that  $X \in MDA(Z)$ .

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**Definition:** (The generalized extreme value distribution (gevd))

Let the distribution function  $H_\gamma$  be given as follows:

$$H_\gamma(x) = \begin{cases} \exp\{-(1 + \gamma x)^{-1/\gamma}\} & \text{wenn } \gamma \neq 0 \\ \exp\{-\exp\{-x\}\} & \text{wenn } \gamma = 0 \end{cases}$$

where  $1 + \gamma x > 0$ , i.e. the definition area of  $H_\gamma$  is given as

$$\begin{aligned} x &> -\gamma^{-1} && \text{wenn } \gamma > 0 \\ x &< -\gamma^{-1} && \text{wenn } \gamma < 0 \\ x &\in \mathbb{R} && \text{wenn } \gamma = 0 \end{aligned}$$

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**Theorem:** (Characterisation of  $MDA(H_\gamma)$ )

- ▶  $F \in MDA(H_\gamma)$  with  $\gamma > 0 \iff F \in MDA(\Phi_\alpha)$  with  $\alpha = 1/\gamma > 0$ .
- ▶  $F \in MDA(H_0) \iff F \in MDA(\Lambda)$ .
- ▶  $F \in MDA(H_\gamma)$  with  $\gamma < 0 \iff F \in MDA(\Psi_\alpha)$  with  $\alpha = -1/\gamma > 0$ .

# MDAs: Examples and Characterisations



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**Observation:**  $\lim_{x \rightarrow +\infty} \frac{\bar{\Phi}_\alpha(x)}{x^{-\alpha}} = 1, \forall \alpha > 0$ . Thus for  $\Phi_\alpha \in MDA(\Phi_\alpha)$  we have  $\bar{\Phi}_\alpha \in RV_{-\alpha}$ . Does this generally hold for members of  $MDA(\Phi_\alpha)$ ?

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**Theorem:** ( $MDA(\Phi_\alpha)$ , Gnedenko 1943)

$F \in MDA(\Phi_\alpha) (\alpha > 0) \iff \bar{F} \in RV_{-\alpha} (\alpha > 0)$ .

If  $F \in MDA(\Phi_\alpha)$ , then  $\lim_{n \rightarrow \infty} a_n^{-1} M_n = \Phi_\alpha$  with  $a_n = F^{\leftarrow}(1 - n^{-1})$ .

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**Examples:** The following distributions belong to  $MDA(\Phi_\alpha)$ :

- ▶ Pareto:  $F(x) = 1 - x^{-\alpha}, x > 1, \alpha > 0$ .
- ▶ Cauchy:  $f(x) = (\pi(1 + x^2))^{-1}, x \in \mathbb{R}$ .
- ▶ Student:  $f(x) = \frac{\Gamma((\alpha+1)/2)}{\sqrt{\alpha\pi}\Gamma(\alpha/2)(1+x^2/\alpha)^{(\alpha+1)/2}}, \alpha \in \mathbb{N}, x \in \mathbb{R}$ .
- ▶ Loggamma:  $f(x) = \frac{\alpha^\beta}{\Gamma(\beta)} (\ln x)^{\beta-1} x^{-\alpha-1}, x > 1, \alpha, \beta > 0$ .

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**Example:** Let  $X \sim U(0, 1)$ . it holds  $X \in MDA(\Psi_1)$  with  $a_n = 1/n$ ,  $n \in \mathbb{N}$ .