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**Simplified case:**  $Y_i$  are independent for i = 1, 2, ..., m. Let  $\Omega = \{0, 1\}^m$  be the state space of the random vector Y. Consider the probability measure P in  $\Omega$ :

$$P(\{y\}) = \prod_{i=1}^{m} ar{p}_{i}^{y_{i}} (1 - ar{p}_{i})^{1 - y_{i}}, \ y \in \{0, 1\}^{m}.$$

The moment generating function of L is  $M_L(t) = \prod_{i=1}^m (e^{te_i}\bar{p}_i + 1 - \bar{p}_i)$ .

$$Q_t(\{y\}) = \prod_{i=1}^n \left( \frac{\exp\{te_i y_i\}}{\exp\{te_i\}\bar{p}_i + 1 - \bar{p}_i} \bar{p}_i^{y_i} (1 - \bar{p}_i)^{1 - y_i} \right).$$

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The solution  $t_z$  (which clearly depends also on the given c) specifies the correct *degree of tilting*.

(2) Generate n<sub>1</sub> conditional realisations of the vector of default indicators (Y<sub>1</sub>,..., Y<sub>m</sub>), Y<sub>i</sub> are simulated from Bernoulli(q<sub>i</sub>), i = 1, 2, ..., m, with

$$q_i = \frac{\exp\{t_z e_i\}p_i(z)}{\exp\{t_z e_i\}p_i(z) + 1 - p_i(z)}$$

(3) Let M<sub>L</sub><sup>(z)</sup>(t) := ∏[exp{t<sub>z</sub>e<sub>i</sub>}p<sub>i</sub>(z) + 1 − p<sub>i</sub>(z)] be the conditional moment generating function of L. Let L<sup>(1)</sup>, L<sup>(2)</sup>,...,L<sup>(n<sub>1</sub>)</sup> be the n<sub>1</sub> conditional realisations of L for the n<sub>1</sub> simulated realisations of Y<sub>1</sub>, Y<sub>2</sub>,..., Y<sub>m</sub>. Compute the *IS*-estimator for the tail probability of the conditional loss distribution:

$$\hat{\theta}_{n_1}^{(IS)}(z) = M_L^{(z)}(t) \frac{1}{n_1} \sum_{j=1}^{n_1} I_{L^{(j)} \ge c} \exp\{-t_z L^{(j)}\} L^{(j)}.$$

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Naive approach: Generate many realisations z of the impact factors Z and compute  $\hat{\theta}_{n_1}^{(IS)}(z)$  for every one of them. The required estimator is the average of  $\hat{\theta}_{n_1}^{(IS)}(z)$  over all realisations z.

This is not the most efficient approach, see Glasserman and Li (2003).

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Assumption:  $Z \sim N_k(0, \Sigma)$  (e.g. like in the probit-normal or logit BMD), where  $k \in \mathbb{N}$  is the number of factors.

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Let the IS density g be the density of  $N_k(\mu, \Sigma)$  for a new expected vector  $\mu \in \mathbb{R}^k$ . A good choice of  $\mu$  should lead to frequent realisations of z which imply high conditional default probabilities  $p_i(z)$ .

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The likelihood ratio:

$$r_{\mu}(Z) = \frac{\exp\{-\frac{1}{2}Z^{t}\Sigma^{-1}Z\}}{\exp\{-\frac{1}{2}(Z-\mu)^{t}\Sigma^{-1}(Z-\mu)\}} = \exp\{-\mu^{t}\Sigma^{-1}Z + \frac{1}{2}\mu^{t}\Sigma^{-1}\mu\}$$

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- (1) Generate  $z_1, z_2, ..., z_n \sim N_k(\mu, \Sigma)$  (*n* is the number of the simulation rounds)
- (2) For each  $z_i$  compute  $\hat{\theta}_{n_1}^{(IS)}(z_i)$  by applying the IS algorithm for the conditional loss.

Assumption:  $Z \sim N_k(0, \Sigma)$  (e.g. like in the probit-normal or logit BMD), where  $k \in \mathbb{N}$  is the number of factors.

Let the IS density g be the density of  $N_k(\mu, \Sigma)$  for a new expected vector  $\mu \in \mathbb{R}^k$ . A good choice of  $\mu$  should lead to frequent realisations of z which imply high conditional default probabilities  $p_i(z)$ .

The likelihood ratio:

$$r_{\mu}(Z) = \frac{\exp\{-\frac{1}{2}Z^{t}\Sigma^{-1}Z\}}{\exp\{-\frac{1}{2}(Z-\mu)^{t}\Sigma^{-1}(Z-\mu)\}} = \exp\{-\mu^{t}\Sigma^{-1}Z + \frac{1}{2}\mu^{t}\Sigma^{-1}\mu\}$$

**Algorithm:** complete IS for Bernoulli mixture models with Gaussian factors

- (1) Generate  $z_1, z_2, \ldots, z_n \sim N_k(\mu, \Sigma)$  (*n* is the number of the simulation rounds)
- (2) For each  $z_i$  compute  $\hat{\theta}_{n_1}^{(IS)}(z_i)$  by applying the IS algorithm for the conditional loss.
- (3) compute the IS estimator for the independent excess probability:

$$\hat{\theta}_n^{(IS)} = \frac{1}{n} \sum_{i=1}^n r_\mu(z_i) \hat{\theta}_{n_1}^{(IS)}(z_i)$$

## The choice of $\boldsymbol{\mu}$

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A sketch of the idea of Glasserman and Li (2003):

Since  $\hat{\theta}_{n_1}^{(IS)}(z) \approx \mathbb{P}(L \ge c | Z = z)$ , search for an appropriate IS density for the function  $z \mapsto \mathbb{P}(L \ge c | Z = z)$ .

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Approach:

a) the IS densiity  $g^*$  should be proportional to  $\mathbb{P}(L \ge c | Z = z) \exp\{-\frac{1}{2}z^t \Sigma^{-1}z\}.$ 

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Approach:

a) the IS density  $g^*$  should be proportional to  $\mathbb{P}(I > a|Z = r) \exp\left(-\frac{1}{2}a^{T}\Sigma^{-1}r\right)$ 

 $\mathbb{P}(L \geq c | Z = z) \exp\{-\frac{1}{2}z^t \Sigma^{-1}z\}.$ 

b) the IS density should be a multivariate normal distribution  $N_k(\mu, \Sigma)$  where the expected vector  $\mu$  is chosen as follows:

$$\mu = \operatorname{argmax}_{z} \left\{ \mathbb{P}(L \ge c | Z = z) \exp\{-\frac{1}{2}z^{t}\Sigma^{-1}z\} \right\}.$$

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This problem is hard to solve exactly; in general  $\mathbb{P}(L \ge c | Z = z)$  is not available in analytical form.

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Approach:

a) the IS density  $g^*$  should be proportional to  $\mathbb{P}(t > t < \tau) = (-1, t < \tau)$ 

 $\mathbb{P}(L \geq c | Z = z) \exp\{-\frac{1}{2}z^t \Sigma^{-1}z\}.$ 

b) the IS density should be a multivariate normal distribution  $N_k(\mu, \Sigma)$  where the expected vector  $\mu$  is chosen as follows:

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Glasserman und Li (2003) propose some numerical solution approaches.

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