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The Clayton copulas have a lower tail dependence.



Example:

Let $\phi(t) = 1 - t$, $t \in [0,1]$. Then $\phi^{[-1]}(t) = \max\{1 - t, 0\}$ and $C_{\phi}(u_1, u_2) := \phi^{[-1]}(\phi(u_1) + \phi(u_2)) = \max\{u_1 + u_2 - 1, 0\} = W(u_1, u_2)$. Thus the Fréchet lower bound is an Archimedian copula.

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Theorem: Let $(X_1, X_2)^T$ be a random vector with continuous marginal distributions and an Archimedian copula C generated by ϕ . Then $\rho_{\tau}(X_1, X_2) = 1 + 4 \int_0^1 \frac{\phi(t)}{\phi'(t)} dt$ holds.

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Definition: A function $g:[0,\infty)\to [0,\infty)$ is called completely monotone iff all higher order derivatives of g exist and the following inequalities hold for $k\in\mathbb{N}$: $(-1)^k\left.\left(\frac{d^k}{ds^k}g(s)\right)\right|_{s=t}\geq 0,\ \forall t\in(0,\infty).$

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Theorem: (Kimberling 1974)

Let $\phi \colon [0,1] \to [0,\infty]$ be a continuous, strictly monotone decreasing function with $\phi(0) = \infty$ and $\phi(1) = 0$. The function $C \colon [0,1]^d \to [0,1]$, $C(u) := \phi^{-1}(\phi(u_1) + \phi(u_2) + \ldots + \phi(u_d))$ is a copula for $d \ge 2$ iff ϕ^{-1} is completely monotone on $[0,\infty)$.

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Lemma: A function $\psi \colon [0,\infty) \to [0,\infty)$ is completely monotone with $\psi(0)=1$ iff ψ is the Laplace-Stieltjes transform of some distribution function G on $[0,\infty)$, i.e. $\psi(s)=\int_0^\infty e^{-sx}dG(x)$, $s\geq 0$.

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Advantages and disadvantages of Archimedian copulas:

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and the distribution function of $U = (U_1, U_2, \dots, U_d)$ is an Archimedian copula with generator ψ^{-1} .

Advantages and disadvantages of Archimedian copulas:

- can model a broader class of dependencies
- have a closed form representation

Theorem: Let G be a distribution function on $[0,\infty)$ such that G(0)=0. Let ψ be the Laplace-Stieltjes transform of G, i.e. $\psi(s)=\int_0^\infty e^{-sx}dG(x)$ for $s\geq 0$. Let X be a r.v. with distribution function G and let U_1,U_2,\ldots,U_d be conditionally independent r.v. for $X=x,\,x\in{\rm I\!R}^+$, with conditional distribution function $F_{U_k|X=x}(u)=\exp(-x\psi^{-1}(u))$ for $u\in[0,1]$.

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Multivariate Archimedian copulas (contd.)

Theorem: Let G be a distribution function on $[0,\infty)$ such that G(0)=0. Let ψ be the Laplace-Stieltjes transform of G, i.e. $\psi(s)=\int_0^\infty e^{-sx}dG(x)$ for $s\geq 0$. Let X be a r.v. with distribution function G and let U_1,U_2,\ldots,U_d be conditionally independent r.v. for $X=x,\,x\in {\rm I\!R}^+$, with conditional distribution function $F_{U_k|X=x}(u)=\exp(-x\psi^{-1}(u))$ for $u\in[0,1]$.

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and the distribution function of $U = (U_1, U_2, \dots, U_d)$ is an Archimedian copula with generator ψ^{-1} .

Advantages and disadvantages of Archimedian copulas:

- can model a broader class of dependencies
- have a closed form representation
- depend on a small number of parameters in general
- the generator function needs to fulfill quite restrictive technical assumptions



Observe: Consider a symmetric positive definite matrix $R \in \mathbb{R}^{d \times d}$ and its Cholesky factorization $AA^T = R$ with $A \in \mathbb{R}^{d \times d}$. If $Z_1, Z_2, \ldots, Z_d \sim N(0,1)$ are independent, then $\mu + AZ \sim N_d(\mu, R)$.

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Algorithm: for the generation of a random vector $U = (U_1, U_2, \dots, U_d)$ whose distribution function is the copula C_R^{Ga} , R positive definite with all ones on the main diagonal.

Compute the Cholesly factorization of $R: R = AA^T$.

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- ▶ Compute the Cholesly factorization of R: $R = AA^T$.
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- ▶ Output $U = (U_1, U_2, ..., U_d)$; U has distribution function C_R^{Ga} .

Algorithm: for the generation of a random vector $U = (U_1, U_2, \dots, U_d)$ whose distribution function is the copula $C_{\nu,R}^t$, R positive definite with all ones on the main diagonal, $\nu \in \mathbb{N}$.

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The generator $\varphi(t)=(t^{-\theta}-1)/\theta,\, \theta>0$ yields the Clayton copula C_{θ}^{Cl} .



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For
$$X \sim Gamma(1/\theta, 1)$$
 with d.f. $f_X(x) = (x^{1/\theta - 1}e^{-x})/\Gamma(1/\theta)$ we have: $E(e^{-sX}) = \int_0^\infty e^{-sx} \frac{1}{\Gamma(1/\theta)} x^{1/\theta - 1} e^{-x} dx = (s+1)^{-1/\theta} = \tilde{\varphi}^{-1}(s)$.

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Simulation of the Gumbel copula ($\theta \ge 1$)

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Let X be a positive stable r.v., $X \sim St(1/\theta,1,\gamma,0)$ with $\gamma = (\cos(\pi/(2\theta)))^{\theta} > 0$ (and $\alpha = \frac{1}{\theta}$, $\beta = 1$, $\delta = 0$)

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Set
$$(Z_1, Z_2)^T := (VS^{\theta}, (1 - V)S^{\theta})^T$$
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The distribution function of $(\bar{F}(Z_1), \bar{F}(Z_2))^T$ is C_{θ}^{Gu} . Convince yourself!

Algorithm to generate a random vector $U = (U_1, U_2, \dots, U_d)$ with the Gumbel copula C_{θ}^{Gu} as distribution function.

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- ► Simulate two i.i.d. r.v. V_1 , $V_2 \sim U(0,1)$.
- Simulate two independent r.v. W_1 , W_2 with $W_1 \sim \Gamma(1,1)$, $W_2 \sim \Gamma(2,1)$
- Set $S := I_{V_2 < 1/\theta} W_1 + I_{V_2 > 1/\theta} W_2$.
- Set $(Z_1, Z_2) := (V_1 S^{\theta}, (1 V_1) S^{\theta}).$
- ▶ The distribution function of $U = \left(\exp(-Z_1^{1/\theta}), \exp(-Z_2^{1/\theta}) \right)^T$ is C_{θ}^{Gu} .

Algorithm to generate a random vector $U = (U_1, U_2, \dots, U_d)$ with the Gumbel copula C_{θ}^{Gu} as distribution function.

Input: The dimension $d \in \mathbb{N}$, the parameter $\theta \geq 1$.

- ► Simulate two i.i.d. r.v. V_1 , $V_2 \sim U(0,1)$.
- Simulate two independent r.v. W_1 , W_2 with $W_1 \sim \Gamma(1,1)$, $W_2 \sim \Gamma(2,1)$
- Set $S := I_{V_2 < 1/\theta} W_1 + I_{V_2 > 1/\theta} W_2$.
- Set $(Z_1, Z_2) := (V_1 S^{\theta}, (1 V_1) S^{\theta}).$
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Question 2: Estimation of the parameters of the prespecified family of copulas used for the modelling?